

On the free energy of a directed polymer in a Brownian environment

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Abstract

We prove a formula conjectured in [14] for the free energy density of a directed polymer in a Brownian environment in $1 + 1$ dimensions.

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1 Introduction

Let $B^{(1)}, B^{(2)}, \dots$ be independent standard one-dimensional Brownian motions. Denote the increments of $B^{(i)}$ by $B_{(s,t)}^{(i)} = B_t^{(i)} - B_s^{(i)}$.

For $\beta \in \mathbb{R}$ set

$$Z_n(\beta) = \int_{0 < s_1 < \dots < s_{n-1} < n} ds_1 \dots ds_{n-1} \exp \{ \beta (B_{(0,s_1)}^{(1)} + \dots + B_{(s_{n-1},n)}^{(n)}) \}. \quad (1)$$

This is the partition function for a continuous model of a directed polymer in a Brownian environment in $1 + 1$ dimensions. In the paper [14], using queueing-theoretic ideas in the context of geometric functionals of Brownian motion, certain limiting results were obtained which led the authors to conjecture an explicit formula for the free energy density

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log Z_n(\beta), \quad (2)$$

namely that it should be given by, almost surely,

$$f(\beta) = \begin{cases} -(-\Psi)^*(-\beta^2) - 2 \log |\beta| & : \beta \neq 0 \\ 1 & : \beta = 0 \end{cases} \quad (3)$$

where $\Psi(m) \equiv \Gamma'(m)/\Gamma(m)$ is the restriction of the digamma function to $(0, \infty)$, and $(-\Psi)^*$ is the convex dual of the function $-\Psi$. The aim of this paper is to give a rigorous proof of this conjecture. The proof uses tools

from large deviation theory. As a corollary we give a new proof that $c \geq 2$, where

$$\lim_{n \rightarrow \infty} \frac{1}{n} L_n(n) = c \quad \text{a.s.} \quad (4)$$

and

$$L_n(t) = \sup_{0 \leq s_1 \leq \dots \leq s_{n-1} \leq t} B_{(0, s_1)}^{(1)} + \dots + B_{(s_{n-1}, t)}^{(n)}.$$

It was proved using direct methods in [11] that $c = 2$, where the authors also describe how the result may be deduced from the theory of random matrices.

The directed polymer model we have discussed here, and for which we have computed the free energy density, is a continuous version of the classical two-dimensional directed polymer, where it is not known how to compute the free energy density (see, for example, Derrida [6], Carmona and Hu [4]). Recent work on a continuous model different to this can be found in Comets and Yoshida [5].

In the next section we recall the framework which was developed in [14] to extend some standard constructions from queueing theory to the context of geometric functionals of Brownian motion, and explain how this leads to the conjectured formula for the free energy density. Section 3 is devoted to the proof of the main result. In section 4 we prove that f is analytic and strictly convex, and record a large deviation principle that we will use in section 5 to prove (4).

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2 Generalised Brownian queues

In this section we recall the framework which was developed in [14] to extend some standard constructions from queueing theory to the context of geometric functionals of Brownian motion, and explain how this leads to the conjectured formula for the free energy density.

The *generalised Brownian queue* is characterised as follows. Let B and C be two independent standard Brownian motions indexed by the entire real line, and write

$$B_{(s,t)} = B_t - B_s, \quad C_{(s,t)} = C_t - C_s$$

Fix $m > 0$ and, for $t \in \mathbb{R}$, set

$$\begin{aligned} r(t) &= \log \int_{-\infty}^t ds \exp\{B_{(s,t)} + C_{(s,t)} - m(t-s)\} \\ f(s,t) &= B_{(s,t)} + r(s) - r(t) \\ g(s,t) &= C_{(s,t)} + r(s) - r(t) \end{aligned}$$

and define $f : \mathbb{R} \rightarrow \mathbb{R}$ by $f(t) = f(0, t)$.

To put this in context, the ‘Brownian queue’ is defined similarly but with ‘ $\log \int \exp$ ’ replaced by ‘ \sup ’ (see, for example, [14]). The usual $M/M/1$ queue is defined similarly to the Brownian queue but with Brownian motions replaced by Poisson counting processes. Thus the Brownian motions B_t and $mt - C_t$ can be thought of, respectively, as the *arrivals* and *service* processes, r as the *queue-length process* and f as the *output*, or *departure*, process.

In [14] it is shown, using results of Matsumoto and Yor [12, 13], that the generalised Brownian queue is *quasi-reversible*, that is: f is a standard Brownian motion, and $\{f(s), s \leq t\}$ is independent of $r(t)$. We can thus consider a sequence of generalised Brownian queues in tandem and expect this ‘queueing network’ to have nice properties (analogous to the ‘product-form solutions’ of classical queueing theory).

Let $B, B^{(1)}, B^{(2)}, \dots$ be a sequence of independent standard Brownian motions, each indexed by \mathbb{R} , and let $m > 0$ be a fixed constant. For $-\infty < s \leq t < \infty$, set

$$\begin{aligned} r_1(t) &= \log \int_{-\infty}^t ds \exp\{B_{(s,t)} + B_{(s,t)}^{(1)} - m(t-s)\} \\ f_1(s,t) &= B_{(s,t)} + r_1(s) - r_1(t) \\ g(s,t) &= B_{(s,t)}^{(1)} + r_1(s) - r_1(t), \end{aligned}$$

for each $k = 2, 3, \dots$ set

$$\begin{aligned} r_k(t) &= \log \int_{-\infty}^t ds \exp\{f_{k-1}(s,t) + B_{(s,t)}^{(k)} - m(t-s)\} \\ f_k(s,t) &= f_{k-1}(s,t) + r_k(s) - r_k(t), \end{aligned}$$

and for all k define $f_k : \mathbb{R} \rightarrow \mathbb{R}$ by $f_k(t) = f_k(0, t)$.

Note that $r_1(t)$ is clearly stationary in t ; to see that $r_1(0) < \infty$ almost surely simply note that, with probability one, $B_{(s,0)} + B_{(s,0)}^{(1)} + ms < ms/2$ for all s sufficiently negative (by Strassen’s law of the iterated logarithm, for example). In fact, $r_1(0)$ has the same law as $-\log Z_m$, where Z_m is gamma-distributed with parameter m : this is Dufresne’s identity [8, 7].

We first state the quasi-reversibility property, as presented in [14].

Theorem 1 *1. f_1 and g are independent standard Brownian motions indexed by \mathbb{R} .*

2. For each $t \in \mathbb{R}$, $\{(f_1(s), g(s)), -\infty < s \leq t\}$ is independent of $\{r_1(s), s \geq t\}$.

It follows from Theorem 1 that $r_1(0), r_2(0), \dots$ is a sequence of i.i.d. random variables, each distributed as $-\log Z_m$. By construction, we have

$$\sum_{k=1}^n r_k(0) = \log \left[\int_{-\infty}^0 du \exp(B_{(u,0)} + mu) \right. \\ \left. \times \int_{u < s_1 < \dots < s_{n-1} < 0} ds_1 \dots ds_{n-1} \exp\{B_{(u,s_1)}^{(1)} + \dots + B_{(s_{n-1},0)}^{(n)}\} \right] \quad (5)$$

Applying the strong law of large numbers, it can be deduced (see [14] for details) that:

Theorem 2 For each $m > 0$:

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \int_{-\infty < u < s_1 < \dots < s_{n-1} < 0} du ds_1 \dots ds_{n-1} \\ \exp\{mu + B_{(u,s_1)}^{(1)} + \dots + B_{(s_{n-1},0)}^{(n)}\} = -\Psi(m)$$

almost surely, where

$$\Psi(m) = E \log Z_m = \Gamma'(m)/\Gamma(m)$$

is the digamma function (and Γ is the Gamma function).

Theorem 2 can be interpreted as follows. Let \mathcal{B} denote the σ -field generated by the Brownian motions $B^{(1)}, B^{(2)}, \dots$, and let τ_1, τ_2, \dots be the points of a unit-rate Poisson process on \mathbb{R}_+ , independent of \mathcal{B} . For $t_0, t_1, \dots, t_n \in \mathbb{R}$ define

$$\begin{aligned} E_n(t_0, t_1, \dots, t_n) &= B_{(t_0, t_1)}^{(1)} + \dots + B_{(t_{n-1}, t_n)}^{(n)} \\ F_n^m(t_0, t_1, \dots, t_{n-m+1}) &= \exp(B_{(t_0, t_1)}^{(n)} + \dots + B_{(t_{n-m}, t_{n-m+1})}^{(m)}) \\ F_n &= F_n^1 \end{aligned}$$

By Brownian scaling, Theorem 2 is equivalent to:

Theorem 3 For $\theta \neq 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log E[\exp(\theta E_n(0, \tau_1, \dots, \tau_n)) | \mathcal{B}] = -2 \log |\theta| - \Psi(1/\theta^2),$$

almost surely.

Thus, if we set

$$\Lambda(\theta) = \begin{cases} -2 \log |\theta| - \Psi(1/\theta^2), & \theta \neq 0 \\ 0, & \theta = 0, \end{cases}$$

we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log E[\exp(\theta E_n(0, \tau_1, \dots, \tau_n)) | \mathcal{B}] = \Lambda(\theta),$$

almost surely. From the asymptotic expansion

$$\Psi(x) \sim \log x - \frac{1}{2x} - \sum_{k=1}^{\infty} \frac{B_{2k}}{2k x^{2k}} \quad (6)$$

as $x \rightarrow \infty$ (see, for example, Abramowitz and Stegun [1]), we have that Λ is finite and differentiable everywhere, with $\Lambda(0) = \Lambda'(0) = 0$. It follows that the sequence $\frac{1}{n} E_n(0, \tau_1, \dots, \tau_n)$ satisfies the following conditional large deviation principle:

Theorem 4 *Given \mathcal{B} , $\frac{1}{n} E_n(0, \tau_1, \dots, \tau_n)$ satisfies a large deviation principle with good rate function*

$$\Lambda^*(x) = \sup_{\theta \in \mathbb{R}} [x\theta - \Lambda(\theta)]$$

almost surely.

This is a *quenched large deviation principle*, associated with the conditional law of large numbers. For example, Theorem 4 implies that given \mathcal{B} , $\frac{1}{n} E_n(0, \tau_1, \dots, \tau_n) \rightarrow 0$ almost surely. Another implication is that for any $x > 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log P(E_n(0, \tau_1, \dots, \tau_n) > xn | \mathcal{B}) = -\Lambda^*(x)$$

almost surely. For two other related large deviation principles see [14].

We will now describe how this relates to the Brownian directed polymer model. It is shown in Lemma 6 using Kingman's subadditive ergodic theorem that there exists a function $\gamma : \mathbb{R} \rightarrow \mathbb{R}$ such that given $x < 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \int_{xn < s_1 < \dots < s_{n-1} < 0} F_n(xn, s_1, \dots, s_{n-1}, 0) ds_1 \dots ds_{n-1} = \gamma(x) \quad (7)$$

almost surely, and it is shown in Lemma 9 that γ is a concave function on $(-\infty, 0)$. Therefore, by (5), Theorem 2 and Laplace's method, we would expect

$$-\Psi(m) = \sup_{x < 0} [mx + \gamma(x)] = (-\gamma)^*(m)$$

and hence by inversion, $\gamma = -(-\Psi)^*$. The free energy density for our model of a directed polymer in a Brownian environment in 1+1 dimensions, defined in (2), can then be expressed in terms of the digamma function by first using the Brownian scaling property:

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{1}{n} \log Z_n(\beta) &= \gamma(-\beta^2) - 2 \log |\beta| \\ &= -(-\Psi)^*(-\beta^2) - 2 \log |\beta|. \end{aligned}$$

The heuristic argument above is made rigorous by the following Theorem, which is the main result of this paper.

Theorem 5 *Almost surely,*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log Z_n(\beta) = f(\beta)$$

where f is defined by (3), $\Psi(m) \equiv \Gamma'(m)/\Gamma(m)$ is the restriction of the digamma function to $(0, \infty)$, and $(-\Psi)^*$ is the convex dual of the function $-\Psi$.

3 Proofs

We begin by defining for $t \geq 0$ and $x < 0$

$$L_n(t) = \sup_{0 \leq s_1 \leq \dots \leq s_{n-1} \leq t} E_n(0, s_1, \dots, s_{n-1}, t) \quad (8)$$

$$k_{m,n}(x) = \log \int_{\substack{xn < s_1 < \dots \\ \dots < s_{n-m} < xm}} F_{n+1}^{m+1}(xn, s_1, \dots, s_{n-m}, xm) ds_1 \dots ds_{n-m}$$

$$\mathcal{Z}_n(x) = \int_{xn < s_1 < \dots < s_{n-1} < 0} F_n(xn, s_1, \dots, s_{n-1}, 0) ds_1 \dots ds_{n-1} \quad (9)$$

$$\gamma_n(x) = \frac{1}{n} \log \mathcal{Z}_n(x). \quad (10)$$

and recording the following lemma:

Lemma 6 *1. There exists a constant $c \in \mathbb{R}$ such that $\frac{1}{n} L_n(nt) \rightarrow c\sqrt{t}$ almost surely and in expectation.*

2. There exists a function $\gamma : (-\infty, 0) \rightarrow \mathbb{R}$ such that given $x < 0$,

$$k_{0,n}(x)/n \rightarrow \gamma(x) \quad (11)$$

almost surely.

3. The function γ is continuous on $(-\infty, 0)$.

4. Given $x < 0$, $\lim_{n \rightarrow \infty} \gamma_n(x) = \gamma(x)$ almost surely.

Proof. The first part of the lemma follows from Brownian scaling and Kingman's subadditive ergodic theorem (see [11] for details). For the second part, observe that

$$\begin{aligned} k_{0,m}(x) + k_{m,n}(x) &= \log \int_{\substack{xn < s_1 < \dots < s_n < 0 \\ s_{n-m} < xm < s_{n-m+1}}} F_{n+1}(xn, s_1, \dots, s_n, 0) ds_1 \dots ds_n \\ &\leq k_{0,n}(x) \end{aligned}$$

and k is therefore superadditive for fixed x . By construction we have the required conditions for Kingman's subadditive ergodic theorem, and so we may define a function $\gamma : (-\infty, 0) \rightarrow \mathbb{R}$ by (11). For the third part we have by Brownian scaling, for $x < 0$ and $\delta > x$

$$\begin{aligned} k_{0,n}(x - \delta) &= \log \int_{\substack{0 < s_1 < \dots \\ \dots < s_n < n}} \exp(\sqrt{-x + \delta} E_{n+1}(0, s_1, \dots, s_n, n)) ds_1 \dots ds_n \\ &\quad + n \log(-x + \delta) \\ &\leq \log \int_{\substack{0 < s_1 < \dots \\ \dots < s_n < n}} \exp(\sqrt{-x} E_{n+1}(0, s_1, \dots, s_n, n)) ds_1 \dots ds_n \\ &\quad + (\sqrt{-x + \delta} - \sqrt{-x}) L_{n+1}(n) + n \log(-x + \delta) \\ &= k_{0,n}(x) + (\sqrt{-x + \delta} - \sqrt{-x}) L_{n+1}(n) \\ &\quad - n \log(-x) + n \log(-x + \delta) \end{aligned} \tag{12}$$

where $=_d$ denotes equality in distribution. Let c be the constant in the first part of this lemma; then by Brownian scaling and Slutsky's Lemma (see for example [9]), $\frac{1}{n} L_{n+1}(n) \Rightarrow c$. By the second part of this lemma and (12),

$$\gamma(x - \delta) \leq \gamma(x) + c(\sqrt{-x + \delta} - \sqrt{-x}) - \log(-x) + \log(-x + \delta)$$

By symmetry,

$$V_n := \inf_{0 \leq s_1 \leq \dots \leq s_{n-1} \leq n} E_n(0, s_1, \dots, s_{n-1}, n) =_d -L_n(n) \tag{13}$$

and so we have similarly

$$\gamma(x - \delta) \geq \gamma(x) - c(\sqrt{-x + \delta} - \sqrt{-x}) - \log(-x) + \log(-x + \delta).$$

So

$$\begin{aligned} |\gamma(x - \delta) - \gamma(x)| &\leq c|\sqrt{-x + \delta} - \sqrt{-x}| \\ &\quad + |\log(-x + \delta) - \log(-x)|. \end{aligned} \tag{14}$$

For the fourth part we have

$$\begin{aligned} k_{0,n}(x) &\leq \log \left(e^{2\bar{B}(n,0,x)} \int_{xn < s_1 < \dots < s_n < 0} F_n(xn, s_2, \dots, s_n, 0) ds_1 \dots ds_n \right) \\ &\leq \log \left(xne^{2\bar{B}(n,0,x)} \int_{xn < s_2 < \dots < s_n < 0} F_n(xn, s_2, \dots, s_n, 0) ds_2 \dots ds_n \right) \\ &= \log \left(xne^{2\bar{B}(n,0,x)} \mathcal{Z}_n(x) \right) \end{aligned}$$

where for $y \leq x \leq 0$,

$$\overline{B}(n, x, y) = \max_{i=n-1, n, n+1} \sup_{yn \leq r < s \leq xn} |B^{(i)}(r, s)|. \quad (15)$$

From Borell's inequality and the Borel-Cantelli Lemma, there exists a null set \mathcal{N} such that on its complement \mathcal{N}^c , $\overline{B}(n, x, y)/n \rightarrow 0$ for all x, y , so

$$\gamma(x) \leq \liminf_{n \rightarrow \infty} \gamma_n(x) \quad \text{a.s.}$$

Now for $\epsilon > 0$,

$$\begin{aligned} k_{0,n}(x + \epsilon) &\geq \log \left(e^{-2\overline{B}(n, 0, x+\epsilon)} \int_{\substack{(x+\epsilon)n < s_1 < xn \\ xn < s_2 < \dots \\ \dots < s_n < 0}} F_n(xn, s_2, \dots, s_n, 0) ds_1 \dots ds_n \right) \\ &= \log \left(\epsilon n e^{-2\overline{B}(n, 0, x+\epsilon)} \mathcal{Z}_n(x) \right) \end{aligned}$$

so

$$\gamma(x + \epsilon) \geq \limsup_{n \rightarrow \infty} \gamma_n(x) \quad \text{a.s.}$$

and the result follows by the third part of this lemma. \square

Let $\mathbb{Q}_+ = \mathbb{Q} \cap (0, \infty)$, $\mathbb{Q}_- = -\mathbb{Q}_+$.

Lemma 7 *There exists a null set \mathcal{M} such that the following statement holds on its complement \mathcal{M}^c :*

$\lim_{n \rightarrow \infty} \gamma_n(x) = \gamma(x)$ and $\liminf_{n \rightarrow \infty} \gamma_n(y) \geq \gamma(x)$ for every $x \in \mathbb{Q}_-$ and $y < x$.

Proof. Let $x \in \mathbb{Q}_-$ and $y < x$. From (9) we have

$$\begin{aligned} \mathcal{Z}_n(y) &\geq \int_{xn < s_1 < \dots < s_{n-1} < 0} F_n(yn, s_1, \dots, s_{n-1}, 0) ds_1 \dots ds_{n-1} \\ &= e^{B_{(yn, xn)}^{(n)}} \mathcal{Z}_n(x). \end{aligned} \quad (16)$$

By Lemma 6 there exists a null set \mathcal{N}_x such that on \mathcal{N}_x^c ,

$$\lim_{n \rightarrow \infty} \gamma_n(x) = \gamma(x).$$

With \mathcal{N} as in the proof of Lemma 6, let $\mathcal{M} = \mathcal{N} \cup \bigcup_{x \in \mathbb{Q}_-} \mathcal{N}_x$. \square

Lemma 8 *Almost surely, $\lim_{n \rightarrow \infty} \gamma_n(x) = \gamma(x)$ for all $x < 0$.*

Proof. Choose $x, y \in \mathbb{Q}_-$ with $y < x$. Then if $z \in [y, x]$,

$$\begin{aligned} \mathcal{Z}_n(z) &= \int_{zn < s_1 < \dots < s_{n-1} < 0} F_n(zn, s_1, \dots, s_{n-1}, 0) ds_1 \dots ds_{n-1} \\ &= I_1(n, z) + I_2(n, z) \end{aligned}$$

where

$$\begin{aligned} I_1(n, z) &= e^{B_{(zn, xn)}^{(n)}} \mathcal{Z}_n(x) \\ I_2(n, z) &= \int_{s_1=zn}^{xn} \int_{\substack{s_1 < \dots \\ \dots < s_{n-1} < 0}} F_{n-1}(s_1, \dots, s_{n-1}, 0) ds_2 \dots ds_{n-1} e^{B_{(zn, s_1)}^{(n)}} ds_1 \end{aligned}$$

Now $I_1(n, z) \leq e^{\bar{B}(n, x, y)} \mathcal{Z}_n(x)$ and

$$I_2(n, z) \leq (x - z) n e^{2\bar{B}(n, x, y)} \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right)$$

therefore

$$\sup_{z \in [y, x]} \mathcal{Z}_n(z) \leq e^{\bar{B}(n, x, y)} \mathcal{Z}_n(x) + (x - y) n e^{2\bar{B}(n, x, y)} \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right). \quad (17)$$

Take \mathcal{M} as in the proof of Lemma 7. Now

$$\mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right) \geq e^{-\bar{B}(n, 0, y)} \mathcal{Z}_{n-1}(y)$$

so $\liminf_{n \rightarrow \infty} \frac{1}{n} \log \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right) \geq \gamma(y)$ on \mathcal{M}^c ; and if $\epsilon \in \mathbb{Q}_-$ then

$$\begin{aligned} \mathcal{Z}_n(y + \epsilon) &\geq \int_{\substack{(y+\epsilon)n < s_1 < yn \\ yn < s_2 < \dots < s_{n-1} < 0}} F_{n-1}(yn, s_2, \dots, s_{n-1}, 0) ds_1 \dots ds_{n-1} \\ &\quad \times e^{-2\bar{B}(n, 0, y+\epsilon)} \\ &= (-\epsilon) n e^{-2\bar{B}(n, 0, y+\epsilon)} \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right) \end{aligned}$$

hence

$$\gamma(y + \epsilon) \geq \limsup_{n \rightarrow \infty} \frac{1}{n} \log \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right)$$

on \mathcal{M}^c , and letting $\epsilon \rightarrow 0$ in \mathbb{Q} gives $\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right) = \gamma(y)$ on \mathcal{M}^c . Then by (17) and the proof of Lemma 7, on \mathcal{M}^c

$$\begin{aligned} \gamma(x) &\leq \liminf_{n \rightarrow \infty} \frac{1}{n} \log \inf_{z \in [y, x]} \mathcal{Z}_n(z) \leq \limsup_{n \rightarrow \infty} \frac{1}{n} \log \sup_{z \in [y, x]} \mathcal{Z}_n(z) \\ &\leq \limsup_{n \rightarrow \infty} \frac{1}{n} \log e^{\bar{B}(n, x, y)} \mathcal{Z}_n(x) \vee \\ &\quad \limsup_{n \rightarrow \infty} \frac{1}{n} \log (x - y) n e^{2\bar{B}(n, x, y)} \mathcal{Z}_{n-1}\left(\frac{ny}{n-1}\right) \\ &= \gamma(x) \vee \gamma(y) = \gamma(y). \end{aligned} \quad (18)$$

The result now follows from (14). \square

Lemma 9 *The function γ is concave.*

Proof. For $x, y < 0$ and $\alpha \in (0, 1)$,

$$\gamma_n(\alpha y + (1 - \alpha)x) \geq \frac{[\alpha n]}{n} G_n + \frac{k_n}{n} H_n \quad (19)$$

where

$$\begin{aligned} G_n &= \frac{1}{[\alpha n]} \log \int_{\substack{(\alpha y + (1 - \alpha)x)n < s_1 < \dots \\ \dots < s_{[\alpha n] - 1} < (1 - \alpha)xn}} F_n^{k_n}((\alpha y + (1 - \alpha)x)n, s_1, \dots \\ &\quad \dots, s_{[\alpha n] - 1}, (1 - \alpha)xn) ds_1 \dots ds_{[\alpha n] - 1} \\ k_n &= n - [\alpha n] + 1 \\ H_n &= \frac{1}{k_n} \log \int_{\substack{(1 - \alpha)xn < s_{[\alpha n]} < \dots \\ \dots < s_{n - 1} < 0}} F_{k_n}((1 - \alpha)xn, s_{[\alpha n]}, \dots \\ &\quad \dots, s_{n - 1}, 0) ds_{[\alpha n]} \dots ds_{n - 1} \end{aligned}$$

Now

$$\begin{aligned} G_n &= {}_d \gamma_{[\alpha n]} \left(\frac{\alpha n}{[\alpha n]} y \right) \\ H_n &= \gamma_{k_n} \left(\frac{(1 - \alpha)n}{k_n} x \right). \end{aligned}$$

Choose $w, u \in \mathbb{Q}_-$ with $w < x < u$, and choose $\epsilon, \delta > 0$. Then $\exists n_0 \in \mathbb{N}$ such that $\forall n \geq n_0$, $w < \frac{(1 - \alpha)n}{k_n} x < u$. Also, by (18) $\exists n_1 \in \mathbb{N}$ such that $\forall n \geq n_1$,

$$P\left(\gamma(u) - \epsilon \leq \frac{1}{n} \log \inf_{z \in [w, u]} \mathcal{Z}_n(z) \leq \frac{1}{n} \log \sup_{z \in [w, u]} \mathcal{Z}_n(z) \leq \gamma(w) + \epsilon\right) > 1 - \delta$$

Therefore if $k_n \geq n_0 \vee n_1$,

$$P(\gamma(u) - \epsilon \leq H_n \leq \gamma(w) + \epsilon) > 1 - \delta$$

and therefore $H_n \Rightarrow \gamma(x)$. Similarly $G_n \Rightarrow \gamma(y)$, and hence by Slutsky's Theorem

$$\frac{[\alpha n]}{n} G_n + \frac{k_n}{n} H_n \Rightarrow \alpha \gamma(y) + (1 - \alpha) \gamma(x)$$

Hence by (19),

$$\gamma(\alpha y + (1 - \alpha)x) \geq \alpha \gamma(y) + (1 - \alpha) \gamma(x).$$

□

Proof of Theorem 5 Choose $m > 0$. Define a probability density function on $(-\infty, 0)$ by

$$\mathbb{K}_n^{(m)}(dx) = \frac{1}{\Xi_n(m)} \exp n(mx + \gamma_n(x)) dx$$

where

$$\Xi_n(m) = \int_{-\infty}^0 \exp n(mx + \gamma_n(x)) dx.$$

In the nomenclature of statistical physics, $\mathbb{K}_n^{(m)}$ is the *Kac density* and $\Xi_n(m)$ is the *grandcanonical partition function*. Theorem 2 says that given $\theta > -m$, the convergences

$$\begin{aligned} \frac{1}{n} \log \Xi_n(m) &\rightarrow -\Psi(m) \\ \frac{1}{n} \log \int_{-\infty}^0 e^{n\theta x} \mathbb{K}_n^{(m)}(dx) &\rightarrow \Psi(m) - \Psi(m + \theta) =: \Lambda_m(\theta) \end{aligned} \quad (20)$$

hold almost surely as $n \rightarrow \infty$. Choose $\epsilon > 0$; then since $\Lambda_m^*(\Lambda'_m(0) + \epsilon)$ and $\Lambda_m^*(\Lambda'_m(0) - \epsilon)$ are strictly positive, we can apply the Chernoff bound to give $\mathbb{K}_n^{(m)}((\Lambda'_m(0) - \epsilon, \Lambda'_m(0) + \epsilon)) \rightarrow 1$ almost surely as $n \rightarrow \infty$. Letting $\epsilon \rightarrow 0$ in \mathbb{Q} gives that $\mathbb{K}_n^{(m)}$ is almost surely concentrated on $\Lambda'_m(0) = -\Psi'(m)$ as $n \rightarrow \infty$. Therefore for any $x \in \mathbb{Q}_-$ and $\epsilon > 0$, using (16) we have

$$\begin{aligned} \Xi_n(m) &\geq \int_{x-\epsilon}^x \exp n(my + \gamma_n(y)) dy \\ &\geq \exp\{n(m(x - \epsilon) + \gamma_n(x))\} \int_{x-\epsilon}^x \exp B_{(yn, xn)}^{(n)} dy \\ &\geq \epsilon \exp\{n(m(x - \epsilon) + \gamma_n(x)) - \overline{B}(n, x, x - \epsilon)\} \end{aligned}$$

Therefore by Lemma 8 and (20),

$$-\Psi(m) \geq m(x - \epsilon) + \gamma(x)$$

and we may let $\epsilon \rightarrow 0$ and appeal to the regularity of Ψ to conclude that for all $x < 0$

$$\gamma(x) \leq \inf_{m \in \mathbb{Q}_+} (-mx - \Psi(m)) = -(-\Psi)^*(x).$$

For the reverse inequality we note that for $x \in \mathbb{Q}_-$ and $\epsilon \in (0, -x)$ we may choose $m > 0$ such that $-\Psi'(m) = x + \epsilon$ (see for example [1, 2]). Then by (16),

$$\begin{aligned} \mathbb{K}_n^{(m)}((x, x + 2\epsilon)) &= \frac{1}{\Xi_n(m)} \int_x^{x+2\epsilon} \exp\{n(my + \gamma_n(y))\} dy \\ &\leq \frac{2\epsilon}{\Xi_n(m)} \exp\{n(m(x + 2\epsilon) + \gamma_n(x)) - \overline{B}(n, x + 2\epsilon, x)\} \end{aligned}$$

Therefore using Lemma 8,

$$0 = \lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{K}_n^{(m)}((x, x + 2\epsilon)) \leq m(x + 2\epsilon) + \gamma(x) + \Psi(m)$$

almost surely, in which case we may let $\epsilon \rightarrow 0$ to conclude that

$$\gamma(x) \geq \inf_{m > 0} (-mx - \Psi(m)) = -(-\Psi)^*(x).$$

Therefore $\gamma(x) = -(-\Psi)^*(x)$. Finally (1), the last part of Lemma 6 and Brownian scaling allow us to conclude that almost surely,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log Z_n(\beta) = \begin{cases} \gamma(-\beta^2) - 2 \log |\beta| & : \beta \neq 0 \\ 1 & : \beta = 0 \end{cases}$$

as required. \square

4 Analyticity of the free energy density and a large deviation principle

Theorem 10 *The function f defined in (3) is analytic and strictly convex on \mathbb{R} , $f'(0) = 0$, and $\lim_{\beta \rightarrow \infty} f(\beta)/\beta = 2$.*

Proof For $x < 0$ we have

$$(-\Psi)^*(x) = \sup_{\theta > 0} [x\theta + \Psi(\theta)]$$

Denoting by Ψ_n the n th derivative of the function Ψ , and noting that Ψ_2 is strictly negative everywhere (see for example [2]), we have

$$(-\Psi)^*(x) = x\Psi_1^{-1}(-x) + \Psi(\Psi_1^{-1}(-x))$$

and since Ψ_1 is an invertible analytic function with nonzero derivative, its inverse is analytic. Therefore f is analytic everywhere except possibly at 0.

To investigate the behaviour of f near 0, let $a = \Psi_1^{-1}(\beta^2)$. Then

$$f(\beta) = a\Psi_1(a) - \Psi(a) - \log \Psi_1(a) \quad (21)$$

Now $a \rightarrow \infty$ as $\beta \rightarrow 0$, and from [1] we have

$$\Psi_1(x) \sim \frac{1}{x} + \frac{1}{2x^2} + \frac{1}{6x^3} \quad (x \rightarrow \infty) \quad (22)$$

therefore recalling (6),

$$f(\beta) = 1 + O(a^{-1}) \quad (\beta \rightarrow 0)$$

One further application of (22) to this expression gives $f'(0) = 0$.

Since $f - 1$ is an asymptotic logarithmic moment generating function (see the proof of Lemma 11 below), f is convex. The regularity of Ψ implies further that f is strictly convex.

The power series

$$\log \Gamma(1+z) = -\log(1+z) + z(1-\xi) + \sum_{n=2}^{\infty} (-1)^n [\zeta(n) - 1] z^n / n$$

where ξ is Euler's constant and ζ is the Riemann Zeta Function, is valid for $|z| < 2$ [1]; therefore, since $\Psi(z) = \frac{d}{dz} \log \Gamma(z)$, it may be differentiated to give power series for Ψ and Ψ_1 in a neighbourhood of the origin. Substituting in (21) and letting $a \rightarrow 0$, we obtain $\lim_{\beta \rightarrow \infty} f(\beta)/\beta = 2$. \square

Lemma 11 *Let $\tau_1 \leq \dots \leq \tau_{n-1}$ be the order statistics for $n-1$ independent random variables having the uniform distribution on the interval $[0, n]$. Almost surely, conditional on \mathcal{B} , the random variable $\frac{1}{n} E_n(0, \tau_1, \dots, \tau_{n-1}, n)$ satisfies a large deviation principle with rate function $(f-1)^*$.*

Proof. Choose $\beta \in \mathbb{R}$; then by (1)

$$E[\exp(\beta E_n(0, \tau_1, \dots, \tau_{n-1}, n)) | \mathcal{B}] = \frac{(n-1)!}{n^{n-1}} Z_n(\beta)$$

hence by Stirling's formula and Theorem 5,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log E[\exp(\beta E_n(0, \tau_1, \dots, \tau_{n-1}, n)) | \mathcal{B}] = f(\beta) - 1 \quad (23)$$

almost surely. Now if $\alpha < \nu$ then $Z_n(\alpha) \leq e^{-(\nu-\alpha)V_n} Z_n(\nu)$, where V_n was defined in (13); hence for $\alpha, \beta \in \mathbb{Q}$ with $\beta > \alpha$

$$\begin{aligned} f(\alpha) &\leq \liminf_{n \rightarrow \infty} \frac{1}{n} \log \inf_{\nu \in (\alpha, \beta)} Z_n(\nu) + (\beta - \alpha)c \\ &\leq \limsup_{n \rightarrow \infty} \frac{1}{n} \log \sup_{\nu \in (\alpha, \beta)} Z_n(\nu) + (\beta - \alpha)c \leq f(\beta) + 2(\beta - \alpha)c \end{aligned}$$

almost surely, where c was defined in Lemma 6. Therefore by the continuity of f , there exists a null set \mathcal{N} such that on \mathcal{N}^c , the convergence in (23) holds for all $\beta \in \mathbb{R}$. \square

5 Connection with random matrices and a Brownian directed percolation problem

It was shown in [11] that with L_n defined as in (8), for each $t \geq 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} L_n(nt) = 2\sqrt{t}$$

almost surely. In the notation of Lemma 6 this says that $c = 2$. This result can also be deduced from random matrix theory using the fact [3, 10] that $L_n(1)$ has the same law as the largest eigenvalue of a $n \times n$ GUE random matrix. We note here that the present work allows us to deduce the inequality $c \geq 2$. Since $L_n(n) \geq E_n(0, t_1, \dots, t_{n-1}, n)$ we have that given β ,

$$\frac{1}{n}L_n(n) \geq \frac{1}{\beta} \frac{1}{n} \log E[\exp(\beta E_n(0, \tau_1, \dots, \tau_{n-1}, n)) | \mathcal{B}]$$

almost surely. Letting $n \rightarrow \infty$ and then $\beta \rightarrow \infty$, using (23) and Theorem 10 gives the required inequality.

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